BMJ Global Health

Precision shielding for COVID-19: metrics of assessment and feasibility of deployment

John P A Ioannidis (iD)

Department of Medicine and Department of Epidemiology and Population Health, Stanford University School of Medicine, Stanford, California, USA

Correspondence to
Dr John P A Ioannidis;
jioannid@stanford.edu

ABSTRACT

The ability to preferentially protect high-risk groups in COVID-19 is hotly debated. Here, the aim is to present simple metrics of such precision shielding of people at high risk of death after infection by SARS-CoV-2; demonstrate how they can estimated; and examine whether precision shielding was successfully achieved in the first COVID-19 wave. The shielding ratio, S, is defined as the ratio of prevalence of infection among people in a high-risk group versus among people in a low-risk group. The contrasted risk groups examined here are according to age ($\geq$70 vs <70 years), and institutionalised (nursing home) setting. For age-related precision shielding, data were used from large seroprevalence studies with separate prevalence data for elderly versus non-elderly and with at least 1000 assessed people$\geq$70 years old. For setting-related precision shielding, data were analysed from 10 countries where information was available on numbers of nursing home residents, proportion of nursing home residents among COVID-19 deaths and overall population infection fatality rate (IFR). Across 17 seroprevalence studies, the shielding ratio S for elderly versus non-elderly varied between 0.4 (substantial shielding) and 1.6 (substantial inverse protection, that is, low-risk people being protected more than high-risk people). Five studies in the USA all yielded S=0.4–0.8, consistent with some shielding being achieved, while two studies in China yielded S=1.5–1.6, consistent with inverse protection. Assuming 25% IFR among nursing home residents, S values for nursing home residents ranged from 0.07 to 3.1. The best shielding was seen in South Korea (S=0.07) and modest shielding was achieved in Israel, Slovenia, Germany and Denmark. No shielding was achieved in Hungary and Sweden. In Belgium (S=1.9), the UK (S=2.2) and Spain (S=3.1), nursing home residents were far more frequently infected than the rest of the population. In conclusion, the experience from the first wave of COVID-19 suggests that different locations and settings varied markedly in the extent to which they protected high-risk groups. Both effective precision shielding and detrimental inverse protection can happen in real-life circumstances. COVID-19 interventions should seek to achieve maximal precision shielding.

INTRODUCTION

A major tension in the scientific community regarding the management of the COVID-19

Summary box

- There is heated debate on whether targeted protection of high-risk groups is a feasible strategy for management of the COVID-19 pandemic.
- There is strong evidence for risk stratification in COVID-19 risk of death after infection, with major effects of age, institutionalised setting and other factors.
- It is less known whether high-risk groups defined according to age and institutionalisation criteria have been preferentially protected at all until now in the real-world experience of the first wave of COVID-19.
- The shielding ratio is introduced, as the ratio of the proportion of people infected in a group at high risk of death when infected versus the group at low risk of death when infected.
- Shielding ratios estimated for 17 large seroprevalence studies worldwide show wide variation (0.4–1.6), suggesting a range from substantial shielding to inverse protection (where low-risk people have been more protected than the high-risk ones).
- Shielding ratios calculated for nursing home residents in 10 countries are found to have even greater variation (0.07–3.1).
- Given that both major precision shielding and inverse protection can be seen in real-world data, efforts should be focused at maximising precision shielding and avoiding inverse protection.

---

pandemic is between proponents of targeted approaches, where people at high-risk are preferentially protected, and those who believe that such approaches are practically infeasible.[1–3] The term precision shielding will be used henceforth to denote the extent to which people at higher risk of death (if infected) can be made to be less frequently infected than people for whom infection would carry a lower risk of death.

The tension between these opposing schools may be exacerbated because to-date there is mostly theoretical polarised debate without solid quantification of precision shielding. It would be useful to have standard metrics that can assess whether precision shielding is

achieved. This would allow to explore the feasibility of these debated approaches in real-life data and to monitor the impact of different non-pharmaceutical interventions for COVID-19 using such metrics.

The aim of this article is to present simple metrics of precision shielding; demonstrate how they can be estimated from stratified population seroprevalence data or from information on proportion of deaths occurring in high-risk groups; and examine whether precision shielding was successfully achieved in the first wave of COVID-19, or, conversely, high-risk groups were more frequently infected than low-risk groups ('inverse protection').

CONCEPTUAL BACKGROUND OF PRECISION SHIELDING

Precision shielding stems from the concept of precision medicine and precision public health.[4] The terms stratified medicine, individualised medicine and personalised medicine are also used. The success of the concept has two prerequisites: first, the ability to identify and separate reliably individuals who have very different risks; and second, the availability of effective interventions specifically for those at high risk. The proof that these prerequisites have been met is provided by the improved outcomes of these select, high-risk individuals who are targeted precisely.

There is very strong evidence that the risk of severe adverse outcomes and death in SARS-CoV-2 infected individuals shows extreme risk stratification according to age, and additional substantial risk stratification is possible according to gender, socioeconomic and clinical features.[5–8] Different individuals (eg, children vs debilitated elderly people) vary over 1000-fold in their estimated risk of death and other serious outcomes (eg, hospitalisation), if infected. Therefore, since the first prerequisite is met, the main question is whether the second prerequisite can also be met, that is, whether interventions exist that can offer enhanced protection from SARS-CoV-2 infection targeted to those individuals who are at high risk.

METRICS OF PRECISION SHIELDING

To answer this question, it is important to have some robust metrics that can assess reliably whether precision shielding is achieved or not. The most direct measure is the ratio of prevalence of ensuing infections among people at a high-risk group versus the prevalence of infections among people in a low-risk group. Let us call this ratio, S, the shielding ratio. The contrasted risk groups need to be specified: for example, according to age (eg, ≥70 vs <70 years old), setting (eg, institutionalised vs non-institutionalised), socioeconomic background or multivariable risk scores.

The potential benefit of precision shielding would be greater, when the shielding ratio is lower. A shielding ratio of S=1 means that low-risk and high-risk people are equally frequently infected, a shielding ratio of S=0.5

means that high-risk people have half the risk of being infected than low-risk ones. S may also take values above 1, if somehow high-risk people get more frequently infected than low-risk people, a situation of "inverse protection".

In this framework, the number of lives saved by precision shielding of some high-risk group is proportional to the infection fatality rate (IFR) in the high-risk group, IFR_h, the proportion of the high-risk group in the population, f_h, and the complement of the shielding ratio, $1-S$. The number of COVID-19 deaths in the high-risk group is

$$D_h = S * P * IFR_h * f_h * N \qquad (1)$$

and the number of COVID-19 deaths in the low-risk group is

$$D_l = P * IFR_l * (1 - f_h) * N \qquad (2)$$

where P is the prevalence of the infection in the low-risk group, IFN_l is the infection fatality rate in the low-risk group and N is the total population of interest. The proportion of COVID-19 deaths occurring in the high-risk group is then given by

$$G = D_h / (D_h + D_l) = S * IFR_h * f_h / ((S * IFR_h * f_h) + IFR_l * (1 - f_h)) \qquad (3)$$

ESTIMATION OF PRECISION SHIELDING FROM PREVALENCE DATA

Seroprevalence studies that assess the frequency of infection in the general population (or in samples that try to approximate the general population) can be used to examine the extent of precision shielding achieved in different risk groups. Seroprevalence studies do have several caveats and biases that have been previously described, for example, related to sampling and selection biases and test performance.[9 10] When their data are used to evaluate whether precision shielding has been achieved, some additional issues should be considered. Specifically, the sampling and selection biases and the test performance of the antibody assays may be different in groups of high-risk versus low-risk individuals.

For example, if the high-risk versus low-risk groups are defined based on age, sampling bias and selection forces may be different in the two groups. For example, individuals in poorer health (and thus at higher risk of poor outcomes if infected with SARS-CoV-2) may be less likely to be sampled. This bias may be more prominent in the elderly group in some studies or in the non-elderly group in others. Or, antibody test performance (sensitivity and specificity) may be different in the elderly versus non-elderly group. Some people who are infected do not necessarily seroconvert and this may differ between high-risk and low-risk groups, for example, children may have greater mucosal immunity than older subjects, and therefore have a reduced probability of becoming seropositive when infected. Furthermore, seroreversion rates may also differ.[11]

Acknowledging these caveats, one can assess the shielding ratio S for elderly (eg, ≥70 years old) versus non-elderly in large seroprevalence studies that have substantial amounts of data for both these high-risk and low-risk groups. Here, seroprevalence studies from four recent systematic reviews[9][12–14] were screened. Studies were selected for analysis if they had at least 1000 participants ≥70 years old, so that a substantial amount of data from this group would be available for a meaningful assessment against the younger age group.

Information was extracted on the crude seroprevalence in the elderly versus non-elderly group. Whenever adjusted estimates were provided (eg, adjusting for test performance, demographics, sampling or design features, ethnicity/race and/or other factors), the maximally adjusted estimates were also extracted for each group. Unadjusted and adjusted estimates were compared, but the latter were preferred, whenever available. The default comparison used a 70-year cut-off, but when data were not presented for this cut-off, a lower cut-off (the one most proximal to 70 years) was used. The share of the high-risk group (≥70 years) in the general population was derived from population pyramids for the respective countries/locations.

ESTIMATION OF PRECISION SHIELDING FROM PROPORTION OF COVID-19 DEATHS OCCURRING IN THE HIGH-RISK GROUP

For many risk factors other than age, data on the prevalence of the infection in high-risk and low-risk groups may not be available. In these cases, one can estimate the shielding ratio S if the proportion of COVID-19 deaths that are contributed by the high-risk group (G), the relative share of the high-risk group in the general population (f_h), and the IFR in the low-risk and high-risk groups (IFR_l and IFR_h) are known or can be reasonably approximated. S is then obtained by solving equation 3 for S.

$$S = G * IFR_l * (1 - f_h) / (IFR_h * (f_h - G * f_h)) \qquad (4)$$

For example, extremely high risk of COVID-19 death is seen in institutionalised elderly individuals in nursing homes, where IFR_h is in the range of 25%,[15] that is, about 100-fold higher than in non-institutionalised people. As of this writing, there are no large published seroprevalence studies to-date that have evaluated representative samples of nursing home residents at a national level in different countries. However, one can estimate S by using available data from countries where there is information on the proportion of COVID-19 deaths that happened among institutionalised elderly people, the share of this group in the general population and the IFR in the general population (and hence also in the non-institutionalised population). A range of values 15%–35% for IFR_h may be used in sensitivity analyses with IFR=25% in the baseline scenario.

For the purposes of calculations done here, information of nursing and elderly home beds in different countries and on the proportion of COVID-19 deaths that occurred in nursing homes in the first wave is

derived from the International Long Term Care Policy Network[16] considering the last update of their review on COVID-19 mortality in nursing homes (released on 14 October 2020). Information on the overall population IFR is derived from Ioannidis[9]; data were used only from countries where IFR estimation had been informed from >1500 samples, so that there would be less uncertainty on the IFR estimate. IFR in non-institutionalised people is calculated by the overall IFR excluding the deaths of nursing home residents from the numerator and the number of nursing home residents from the denominator (country population).

PRECISION SHIELDING ACCORDING TO AGE

Table 1 shows the characteristics of 17 eligible large population seroprevalence studies[17–33] where results were separately available in elderly versus younger groups. Table 2 shows the shielding ratios based on crude (unadjusted) and adjusted seroprevalence data. As shown, the shielding ratio ranged from 0.4 to 1.6.

The elderly were infected substantially more frequently than the younger populations in Spain and China and possibly in Hungary, although in Hungary the 95% CI could not exclude values less than 1. Two studies in China yielded very similar estimates of S (1.5–1.6). The Dominican Republic, India and Canada had S values very close to 1.0 (no shielding) and 95% CIs exclude major shielding. Modest shielding was suggested in Brazil (estimated S=0.8), but the 95% CI could not exclude values above 1. In the USA, five studies all show S value estimates below 1, with three of them yielding modest estimates of S=0.8 and the other two yielding S=0.6 and S=0.4, but with largely overlapping CIs. The two lowest estimates come from studies on life insurance applicants and blood donors, where substantial selection biases that depend on age cannot be excluded. For example, stronger healthy volunteer bias may exist in the elderly life insurance applicants and blood donors. The best S values (S=0.5–0.7) suggesting relatively successful age-related shielding were seen in Denmark, Iceland and in two studies in the UK, but the 95% CI had large uncertainty in Denmark.

Of note, most of the seroprevalence estimates in table 1 were from crude, unadjusted analyses (positive samples per total tested samples). Adjustments were used only in the Denmark, Spain, the UK, Brazil, the USA—haemodialysis, USA—New York studies, and adjusted point estimates tended to be similar to the unadjusted ones. The proportion of people ≥70 years old in the general population was 9%–15% in high-income countries and 4%–7% in other countries.

PRECISION SHIELDING ACCORDING TO INSTITUTIONALISED SETTING

Table 3 runs calculations for precision shielding of nursing home residents in 10 countries during the first

Table 1 Characteristics of the eligible seroprevalence studies

Location (reference)	Type of population	Date	Sample tested, n	Age cut-off,* y
Europe				
Denmark (17)	Blood donors	1–26 June	2311	70
Hungary (18)	General	1–16 May	10 472	65
Iceland (19)	General	April–June	30 576	70
Spain (20)	General	14 April to 1 May	61 075	65
UK (21)	General	20 June to 13 July	99 908	65
UK (22)	Biobank	27 May to 14 August	18 734	70
Americas				
Brazil (23)	General	4–7 June	31 165	70
Canada (24)	Blood donors	May	37 737	60
Dominican Republic(25)	General, hotspot areas	April–June	12 897	60
USA (26)	Haemodialysis	July	28 503	65
USA (27)	Life insurance applicants	12 May to 25 June	50 025	61
USA (28)	Blood donors	June–July	189 656	65
USA—New York (29)	General, convenience	19–28 April	15 101	55
USA—Brooklyn (30)	General, convenience	Early May	11 092	70
Asia				
China—Wuchang (31)	General	Mid-May	61 437	70
China (32)	Diverse	January–April	17 368	65
India (33)	General	11 May to 4 June	28 000	60

*For the comparison of seroprevalence in elderly versus younger participants, 70 years was the default cut-off chosen, but as shown in this column, a lower cut-off was chosen when seroprevalence data according to the 70-year cut-off were not available.

wave of COVID-19. They include eight European countries, Israel and South Korea.

As shown, Belgium, UK and Spain have had very unfavourable S values (S=1.9, 2.2. and 3.1 in the baseline scenario, potentially even higher in some sensitivity analyses). This means that people who were institutionalised in nursing homes were approximately two times as likely to be infected than the non-institutionalised population in Belgium and the UK and more than three times more likely to be infected than the non-institutionalised population in Spain.

Sweden and Hungary seemed to have been unable to protect their nursing homes more than the general population and sensitivity analysis for Sweden suggests that it is possible that nursing home residents were even infected more than two times as frequently than the rest of the population. Some substantial shielding was seen in Denmark, Slovenia, Germany and Israel (S=0.3–0.5 in the baseline scenario). Extremely effective shielding was achieved in South Korea, where one can estimate S=0.07. Only 8% of COVID-19 deaths in this country occurred in nursing home residents in the first wave.

PRECISION SHIELDING PLACED INTO CONTEXT

The shielding ratio can be used as a metric to assess whether protection of high-risk populations is being achieved in a given country or jurisdiction. As shown, data from the first wave of COVID-19 suggest that the shielding ratio can take very different values, ranging from extremely effective protection of vulnerable high-risk populations to major inverse protection, where high-risk populations have been protected far less successfully than low-risk populations.

Fatality rates tend to be relatively low in countries where the elderly (and even more so the institutionalised elderly) have been effectively protected. It is possible that one can achieve better values of shielding (lower S) in nursing homes than in non-institutionalised elderly who are unavoidably more freely mobile in the community. Countries that have avoided massive infections in nursing homes have had much lower fatality burden from COVID-19 in the first wave. It is estimated[16] that in the first wave, only 0.01% of South Korean nursing home residents died with COVID-19, as opposed to 3.3% in Sweden and more than 5% in Belgium, England and Spain. While there may be differences on how deaths are attributed to COVID-19 among nursing home residents, these are unlikely to explain away such major differences across countries. Besides nursing homes, some differential protection can be achieved even for the non-institutionalised elderly and this may result in substantially lower fatalities overall. Thus, Iceland and Denmark did have 20% and 35%, respectively, of the COVID-19 deaths occur in nursing homes, but they seem to have protected effectively their comunity-dwelling elderly; therefore, they have had low fatalities in the first wave.

Table 2 Estimates of the shielding ratio for elderly people versus younger people in different locations for the first wave of COVID-19

Location (reference)	Prevalence (%)			Adjusted prevalence (%)			
	Elderly	Younger	S (95% CI)	Elderly	Younger	S (95% CI)	f_h (%)
Europe							
Denmark (17)	1.8	3.0	0.6 (0.4 to 1.1)	1.4	2.5	0.6 (0.3 to 1.1)	14
Hungary (18)	0.8	0.6	1.3 (0.7 to 2.1)	ND	ND	ND	12
Iceland (19)	0.5	1.0	0.5 (0.4 to 0.6)	ND	ND	ND	10
Spain (20)	ND	ND	ND	6.0	4.7	1.3 (1.2 to 1.4)	15
UK (21)	2.7	5.6	0.5 (0.4 to 0.6)	3.3	6.7	0.5 (0.4 to 0.6)	14
UK (22)	6.1	8.8	0.7 (0.6 to 0.8)	ND	ND	ND	14
Americas							
Brazil (23)	2.4	2.8	0.9 (0.7 to 1.1)	2.4	2.8	0.8 (0.6 to 1.2)	6
Canada (24)	0.7	0.7	1.0 (0.7 to 1.3)	ND	ND	ND	12
Dominican Republic (25)	6.0	5.3	1.1 (*)	ND	ND	ND	5
USA (26)	7.6	8.4	0.9 (0.8 to 1.0)	8.1	9.7	0.8 (0.8 to 0.9)	11
USA (27)	2.0	3.1	0.6 (0.5 to 0.8)	ND	ND	ND	11
USA (28)	0.8	1.8	0.4 (0.4 to 0.5)	ND	ND	ND	11
USA—New York (29)	10.9	13.5	0.9 (0.8 to 1.0)	12.1	15.4	0.8 (0.7 to 0.9)	10
USA—Brooklyn (30)	~40	~50	0.8 (*)	ND	ND	ND	9
Asia							
China—Wuchang (31)	3.5	2.2	1.6 (1.4 to 1.8)	ND	ND	ND	7
China (32)	2.0	1.3	1.5 (1.2 to 1.9)	ND	ND	ND	7
India (33)	0.6	0.6	1.0 (0.7 to 1.5)	ND	ND	ND	4

The exact number of participants≥70 years tested is not provided for the studies in Iceland, Canada and Dominical Republic, but based on the number of participants in the highest provided age stratum and the age-structure of the population in these countries, it is likely that those≥70 would exceed the minimum required sample of n=1000. For Iceland, the data include the cases detected with positive test (the majority) plus those estimated to be infected based on antibody testing.

*Information available does not allow reliable calculation of 95% CIs.

f_h, percent share of those ≥70 years old in the population of the location; ND, no data; PRC, polymerase chain reaction; ; S, shielding ratio.

Table 3 Estimating the shielding ratio for institutionalised nursing home residents versus the rest of the population in 10 countries during the first wave of COVID-19

Country	Nursing home residents per 100 000 population	Proportion COVID-19 deaths in nursing homes, G (%)	Overall population IFR (%)	IFR in non-institutionalised, IFR_i (%)	Shielding ratio, S
Belgium	1080	61	0.87	0.34	1.9 (1.4–3.2)
Denmark	690	35	0.27	0.18	0.5 (0.4–0.9)
Germany	980	39	0.23	0.14	0.4 (0.3–0.6)
Hungary	570	23	0.54	0.42	0.9 (0.6–1.4)
Slovenia	1100	81	0.11	0.02	0.3 (0.2–0.5)
Spain	690	63*	0.85	0.31	3.1 (2.2–5.1)
Sweden	810	46	0.57	0.31	1.3 (0.9–2.1)
UK (England)	760	45	0.93	0.51	2.2 (1.6–3.7)
South Korea	420	8	0.09	0.08	0.07 (0.05–0.11)
Israel	520	39	0.10	0.06	0.3 (0.2–0.5)

S is calculated from equation 4 assuming IFR_h in nursing home residents of 25% in the baseline scenario and 15% or 35% in the scenarios of sensitivity analyses that define the presented bounds.

*The proportion may be actually lower, since in its calculation the International Long Term Care Policy Network report[16] uses both confirmed and probable nursing home deaths but only confirmed total deaths; with a proportion of 50% instead of 63%, IFR_i is estimated as 0.43% and S as 2.5. IFR, infection fatality rate.;

The worst fatality rates have been seen in locations with high proportions of elderly and/or institutionalised people and where there was strong inverse protection. For example, Castiglione d'Adda,[34] a small town in Lombardy had 47 COVID-19 deaths in a population of 4550 people. Seroprevalence data[34] showed IgG positivity in 51/155 people≥60 years old versus 64/290 in younger people, which translate into S=1.5 for age-related shielding and the town also had nursing homes affected. Another seroprevalence study in Northern Italy locations found that seroprevalence was 4.5 times larger in nursing home residents compared with non-institutionalised people.[35] While these data may not necessarily be representative of Italy as a whole, they are congruent with the very high fatalities in particular areas of Lombardy in the first wave.[36]

Some countries may have had mixed patterns, for example, protecting somehow their elderly, but not specifically their institutionalised elderly, as in the case of the UK and probably also the USA where 44% of COVID-19 deaths occurred in the 0.59% of the population that resides in nursing homes.[16] This pattern can still translate to heavy cumulative death toll. Institutionalised elderly are at much higher risk of death than other elderly people, and they can contribute a lion's share to the overall death count.

While only age and nursing home residence were explored here, other risk factors may also be assessed in a similar fashion in terms of the extent of precision shielding. For example, socioeconomic factors are known to be strong determinants of the infection rate.[37] Minorities and disadvantaged populations are more likely to be infected and it is possible that may also have more adverse outcomes due to poorer health status.

A research agenda can be built in future work trying to understand correlates and determinants of S. For example, it would be interesting to assess whether S correlates with features of population density, geography, specific non-pharmaceutical interventions and other policies at the population or institutional level (eg, nursing home management, staffing and testing).

Different measures against the spread of COVID-19 need to be assessed in terms of their effect on precision shielding. One might argue that horizontal measures to mitigate COVID-19 for everyone without making discriminations according to risk would have S=1, as infection rates would be decreased equally in all groups. In many/most circumstances, this may not hold true. Most measures may eventually leave some population subgroups more exposed than others. The groups that still remain unavoidably highly exposed may occasionally be among those that have lower risk (eg, young, healthy military personnel in congested areas like barracks or military vessels). However, in most situations horizontal measures may unintentionally leave high-risk groups more exposed than low-risk groups.

For example, horizontal lockdown measures typically protect young, healthy professionals who can work from home, but leave far more exposed the essential workers and those who are disadvantaged, for example, the homeless. These poorly sheltered populations often have a higher burden of background comorbidities and more limited medical care—and are thus at higher risk of death, if infected by SARS-CoV-2. Similarly, horizontal lockdowns may leave nursing home populations less protected than non-institutionalised populations, unless additional targeted measures are taken focused on nursing homes specifically. Nursing home residents have very limited mobility and often live together in closed, congested spaces—as opposed to young, healthy individuals who shelter in place alone or in smaller numbers with their families. Thus, massive infections are easier to occur in nursing homes. The situation may become even worse, if nursing home personnel also has a high S value, since personnel will then infect the residents. This was apparently the case with Stockholm during the first wave, where seroprevalence among nursing home personnel was 23% in the first 20 days of April,[38] three times higher than the general population of Stockholm at the same time. Nursing home personnel in Stockholm was highly mobile and exposed frail elderly across different nursing homes. Lockdown measures also force young low-risk individuals to spend more time indoors and this may increase the exposure of any high-risk family members who have to live in the same house.

CONCLUSIONS

The most-widely used metric for the success of interventions against COVID-19 to-date has been the number of infections. This metric alone is problematic because the vast majority of infections remain unrecorded and the documented infections depend on how many tests are done. A more informative metric of success is the ability of different interventions to generate a most favourable shielding ratio for the most high-risk subgroups of the population. These subgroups may account for the vast majority of the potential deaths and, if properly protected, many deaths can be avoided.

Estimates for the shielding ratio and its evolution as the epidemic wave progresses and as different interventions are employed can be obtained from prevalence studies using antibodies, or other tests (eg, antigen testing).[39 40] Alternatively, they can be obtained from assessments of the profile of fatalities, provided that the relative representation of the high-risk groups of interest in the general population is known, and that some fair estimate of the IFR in low-risk and high-risk groups exists from previous studies. S can be used as an outcome in COVID-19 interventional studies.[41]

Estimates of precision shielding may have some considerable error margin and biases, inherent in the parameters that go in the calculation of S. Thus, one has to interpret the ensuing S values with caution and allow for substantial uncertainty. These values should be able to convey whether some substantial shielding is achieved, or

whether gross inverse protection is making things worse. Small differences in S estimates should not be overinterpreted. Validation with multiple, preferably large and unbiased, studies on the same population may help get a better sense of the accuracy and heterogeneity of such estimates.

Most data analysed here came from high-income countries. Limited data from two other countries (India and Dominican Republic) showed estimates of S very close to 1, suggesting no achieved shielding. It is possible that shielding is more difficult to achieve in some resource-poor settings and in congested, highly mixing populations where most people cannot shelter effectively and have limited private space, living in multigenerational families. Similarly, even within the same country, or sublocation, there may be subpopulations that can achieve much better precision shielding than others, due to socioeconomic circumstances and other factors. Job circumstances may be particularly important and further studies should evaluate precision shielding in different high-risk jobs. For example, some data suggest that age-specific precision shielding can be successfully achieved for healthcare, first response, and public safety personnel: S was 0.4 for such personnel based on a 65-year cut-off in Detroit.[42]

Precision approaches had received enthusiastic support before the COVID-19 era as a way to transform medicine and health at large. Hopes (and hype) were fueled in particular by perceived improvements in predictive ability, especially with the advent of -omics.[43 44] However, the discriminating ability of most -omics discoveries had been modest: for example, single gene variants often differentiated risk by less than 1.1-fold and even complex molecular signatures and multigene models often differentiated risk by less than 3-fold between high-risk and low-risk individuals. In this regard, COVID-19 offers a situation where risk discrimination is far better than most previous efforts at materialising precision medicine. If the risk stratification offered by COVID-19 does not suffice for precision purposes, then it is unlikely that the concept of precision medicine can find fruitful applications with major impact across medicine (perhaps with the exception of some rare conditions). At a minimum, it is worth trying to make precision approaches work for COVID-19. Even modest shielding ratios may translate into hundreds of thousands or even millions of lives saved during the multiyear course of the pandemic.[45]

Data availability statement All data relevant to the study are included in the article.

ORCID iD
John P A Ioannidis http://orcid.org/0000-0003-3118-6859

REFERENCES

1. Smith GD, Spiegelhalter D. Shielding from covid-19 should be stratified by risk. *BMJ* 2020;369:m2063.
2. Great Barrington Declaration. Available: https://gbdeclaration.org/ [Accessed 1 Nov 2020].
3. John Snow Memorandum. Available: https://www.johnsnowmemo.com/ [Accessed 1 Nov 2020].
4. Khoury MJ, Iademarco MF, Riley WT. Precision public health for the era of precision medicine. *Am J Prev Med* 2016;50:398–401.
5. Spiegelhalter D. Use of "normal" risk to improve understanding of dangers of covid-19. *BMJ* 2020;370:m3259.
6. Ioannidis JPA, Axfors C, Contopoulos-Ioannidis DG. Population-Level COVID-19 mortality risk for non-elderly individuals overall and for non-elderly individuals without underlying diseases in pandemic epicenters. *Environ Res* 2020;188:109890.
7. Williamson EJ, Walker AJ, Bhaskaran K, *et al*. Factors associated with COVID-19-related death using OpenSAFELY. *Nature* 2020;584:430–6.
8. Grant A. Apparent reductions in COVID-19 case fatality rates reflect changes in average age of those testing positive. *medRxiv* 2020.
9. Ioannidis JPA. Infection fatality rate of COVID-19 inferred from seroprevalence data. *Bull World Health Organ* 2021;99:19–33 https://www.who.int/bulletin/volumes/99/1/20-265892/en/
10. Burgess S, Ponsford MJ, Gill D. Are we underestimating seroprevalence of SARS-CoV-2? *BMJ* 2020;370:m3364.
11. Seow J, Graham C, Merrick B, *et al*. Longitudinal observation and decline of neutralizing antibody responses in the three months following SARS-CoV-2 infection in humans. *Nat Microbiol* 2020;5:1598–607.
12. Franceschi VB, Santos AS, Glaeser AB. Population-based prevalence surveys during the COVID-19 pandemic: a systematic review. *medRxiv* 2020.
13. Chen X, Chen Z, Azman AS. Serological evidence of human infection with SARS-CoV-2: a systematic review and meta-analysis. *medRxiv* 2020.
14. Arora RK, Joseph A, Van Wyk J, *et al*. SeroTracker: a global SARS-CoV-2 seroprevalence dashboard. *Lancet Infect Dis* 2020;S1473-3099(20)30631–9.
15. Arons MM, Hatfield KM, Reddy SC, *et al*. Presymptomatic SARS-CoV-2 infections and transmission in a skilled nursing facility. *N Engl J Med* 2020;382:2081–90.
16. International Long Term Care Policy Network. Mortality associated with COVID-19 in care homes: international evidence. Available: https://ltccovid.org/2020/04/12/mortality-associated-with-covid-19-outbreaks-in-care-homes-early-international-evidence/#:~:text=Based%20on%20the%20data%20gathered,(based%20on%2021%20countries [Accessed 1 Nov 2020].
17. Pedersen OB, Nissen J, Dinh KM, *et al*. SARS-CoV-2 infection fatality rate among elderly retired Danish blood donors - A cross-sectional study. *Clin Infect Dis* 2020:ciaa1627.
18. Merkely B, Szabó AJ, Kosztin A, *et al*. Novel coronavirus epidemic in the Hungarian population, a cross-sectional nationwide survey to support the exit policy in Hungary. *Geroscience* 2020;42:1063–74.
19. Gudbjartsson DF, Norddahl GL, Melsted P, *et al*. Humoral immune response to SARS-CoV-2 in Iceland. *N Engl J Med* 2020;383:1724–34.
20. Pollán M, Pérez-Gómez B, Pastor-Barriuso R, *et al*. Prevalence of SARS-CoV-2 in Spain (ENE-COVID): a nationwide, population-based seroepidemiological study. *Lancet* 2020;396:535–44.
21. Ward H, Atchinson C, Whitaker M. Antibody prevalence for SARS-CoV-2 following the peak of the pandemic in England: REACT2 study in 100,000 adults [preprint]. *medRxiv* 2020.
22. Biobank UK. *Uk Biobank SARS-CoV-2 serology study*, 2020.
23. Hallal PC, Hartwig FP, Horta BL, *et al*. SARS-CoV-2 antibody prevalence in Brazil: results from two successive nationwide serological household surveys. *Lancet Glob Health* 2020;8:e1390–8.

24 Canadian Blood Services. *COVID-19 seroprevalence report*, 2020.
25 Paulino-Ramirez R, Báez AA, Vallejo Degaudenzi A, *et al*. Seroprevalence of specific antibodies against SARS-CoV-2 from hotspot communities in the Dominican Republic. *Am J Trop Med Hyg* 2020;103:2343–6.
26 Anand S, Montez-Rath M, Han J, *et al*. Prevalence of SARS-CoV-2 antibodies in a large nationwide sample of patients on dialysis in the USA: a cross-sectional study. *Lancet* 2020;396:1335–44.
27 Rigatti SJ, Stout R. SARS-CoV-2 antibody prevalence and association with routine laboratory values in a life insurance applicant population. *medRxiv*2020.
28 Vassallo RR, Bravo MD, Dumont LJ. Seroprevalence of antibodies to SARS-CoV-2 in US blood donors. *medRxiv* 2020.
29 Rosenberg ES, Tesoriero JM, Rosenthal EM, *et al*. Cumulative incidence and diagnosis of SARS-CoV-2 infection in New York. *Ann Epidemiol* 2020;48:23–9.
30 Reifer J, Hayum N, Heszkel B, *et al*. SARS-CoV-2 IgG antibody responses in New York City. *Diagn Microbiol Infect Dis* 2020;98:115128.
31 Pan Y, Li X, Yang G, *et al*. Seroprevalence of SARS-CoV-2 immunoglobulin antibodies in Wuhan, China: part of the city-wide massive testing campaign. *Clin Microbiol Infect* 2020;323. doi:10.1016/j.cmi.2020.09.044. [Epub ahead of print: 06 Oct 2020].
32 Xu X, Sun J, Nie S, *et al*. Seroprevalence of immunoglobulin M and G antibodies against SARS-CoV-2 in China. *Nat Med* 2020;26:1193–5.
33 Murhekar MV, Bhatnagar T, Selvaraju S, *et al*. Prevalence of SARS-CoV-2 infection in India: findings from the National serosurvey, May-June 2020. *Indian J Med Res* 2020;152:48–60.
34 Pagani G, Conti F, Giacomelli A, *et al*. Seroprevalence of SARS-CoV-2 significantly varies with age: preliminary results from a mass population screening. *J Infect* 2020;81:e10–12.
35 Vena A, Berruti M, Adessi A, *et al*. Prevalence of antibodies to SARS-CoV-2 in Italian adults and associated risk factors. *J Clin Med* 2020;9:2780.
36 Boccia S, Ricciardi W, Ioannidis JPA. What other countries can learn from Italy during the COVID-19 pandemic. *JAMA Intern Med* 2020;180:927–8.
37 Iacobucci G. Covid-19: increased risk among ethnic minorities is largely due to poverty and social disparities, review finds. *BMJ* 2020;371:m4099.
38 Lindahl JF, Hoffman T, Esmaeilzadeh M, *et al*. High seroprevalence of SARS-CoV-2 in elderly care employees in Sweden. *Infect Ecol Epidemiol* 2020;10:1789036.
39 Mak GC, Cheng PK, Lau SS, *et al*. Evaluation of rapid antigen test for detection of SARS-CoV-2 virus. *J Clin Virol* 2020;129:104500.
40 Nagura-Ikeda M, Imai K, Tabata S, *et al*. Clinical evaluation of self-collected saliva by quantitative reverse transcription-PCR (RT-qPCR), direct RT-qPCR, reverse transcription-loop-mediated isothermal amplification, and a rapid antigen test to diagnose COVID-19. *J Clin Microbiol* 2020;58:e01438–20.
41 Cristea IA, Naudet F, Ioannidis JPA. Preserving equipoise and performing randomised trials for COVID-19 social distancing interventions. *Epidemiol Psychiatr Sci* 2020;29:1–27.
42 Akinbami LJ, Vuong N, Petersen LR, *et al*. SARS-CoV-2 seroprevalence among healthcare, first response, and public safety personnel, Detroit metropolitan area, Michigan, USA, May-June 2020. *Emerg Infect Dis* 2020;26:2863–71.
43 Collins FS, Varmus H. A new initiative on precision medicine. *N Engl J Med* 2015;372:793–5.
44 Bossuyt PMM. The thin line between hope and hype in biomarker research. *JAMA* 2011;305:2229–30.
45 Ioannidis JPA. Global perspective of COVID-19 epidemiology for a full-cycle pandemic. *Eur J Clin Invest* 2020;50:e13421.